

Forecasting With Kernels

Perspective: *Seminar thesis, Bachelor, Master's thesis*

Writing hints: <https://www.overleaf.com/read/hdpkgxgjkbgw>

Problem Description

Suppose the function $f: \mathbb{R}^d \rightarrow \mathbb{R}$ is observed with function values, $(x_i, f_i) \in \mathbb{R}^d \times \mathbb{R}$, $i = 1, \dots, n$. Employing a kernel function $k: \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$, it holds that

$$\hat{f}(x_i) = f_i, \quad i = 1, \dots, n, \quad (1)$$

where

$$\hat{f}(x) := \sum_{j=1}^n k(x, x_j) w_j$$

and provided that the weights w satisfy the linear system of equations

$$\sum_{j=1}^n k(x_i, x_j) w_j = f_i, \quad i = 1, \dots, n.$$

A canonical candidate for the kernel is the Gaussian kernel $k(x, y) := e^{-\|y-x\|^2 / (2\eta^2)}$ (with Euclidean norm $\|\cdot\|$ in \mathbb{R}^d and characteristic length η).

For observations f_i with noise, the kernel estimator for $f(x)$ is

$$\hat{f}(x) = \sum_{j=1}^n k(x, x_j) w_j \quad (2)$$

with

$$\lambda w_i + \sum_{j=1}^n k(x_i, x_j) w_j = f_i, \quad i = 1, \dots, n. \quad (3)$$

The regression parameter λ is chosen to match the noise: $\lambda = 0$ matches the initial situation so that $\hat{f}(x_i) = f_i$ without noise (interpolation, cf. (1)). A value of λ too close to zero leads to *overfitting*, larger values for λ lead to *oversmoothing* (underfitting). λ acts as a regularization parameter, balancing bias and variance to prevent overfitting while maintaining model flexibility.

Task: Forecasting with Delayed Embedding of Dimension d

We shall employ (2) and (3) to a stationary time series $X = (X_t)$ with $X_t \in \mathbb{R}$, $\mathbb{E} X_t = 0$, and constant variance

$$\sigma^2 := \text{var}(X_{t+1} | X_{t-d+1}, \dots, X_t),$$

where $d \in \mathbb{N}$ is fixed.

Consider the observation (realization)

$$X_{2-d}, \dots, X_1, X_2, \dots, X_{n+1} \quad (4)$$

of the time series. In the *sliding window approach*, set

$$x_i := (X_{i-d+1}, X_{i-d+2}, \dots, X_i) \in \mathbb{R}^d \quad \text{and} \quad f_i := X_{i+1} \in \mathbb{R}, \quad i = 1, \dots, n, \quad (5)$$

then $\hat{f}((X_{t-d+1}, \dots, X_t))$ is an *estimate* for the subsequent X_{t+1} ,

$$\hat{f}((X_{t-d+1}, \dots, X_t)) \approx \mathbb{E}[X_{t+1} | X_{t-d+1}, \dots, X_t].$$

1. Discuss and implement the relations (2), (3) and (5) for some adequate, oscillating time series observations (X_t) as in (4) (for example the *nottem* time series data).
2. **Recursive multistep forecasting:** For some chosen starting values $(\hat{X}_1, \dots, \hat{X}_d)$, extend this time series by setting

$$\hat{X}_{t+1} := \hat{f}((\hat{X}_{t-d+1}, \dots, \hat{X}_t)) + \varepsilon_t, \quad t = d, d+1, \dots$$

and *plot* the realization $\hat{X}_1, \hat{X}_2, \dots, \hat{X}_d, \dots, \hat{X}_n, \dots, \hat{X}_N$ with some $N \gg d$; the iid random errors ε_t have 0-mean, and variance corresponding to the residual variance,

$$\text{var } \varepsilon_t \approx \sigma^2 = \text{var}(X_{t+1} | X_{t-d+1}, \dots, X_t).$$

3. Adjust the parameters λ, d, σ and the kernel k (or the parameter in the kernel, η , say) so that the time series \hat{X} *visually* matches the initial observations (4) of X .
4. Try the norm $\|x\|_\Sigma := x^\top \Sigma^{-1} x$ (Mahalanobis distance) instead of the Euclidean norm, where Σ is the covariance matrix of the random vector (X_{i-d+1}, \dots, X_i) . The covariance matrix can be estimated from the observations (4).

Bonne chance!